

# Some facts about expectation

October 22, 2010

## 1 Expectation identities

There are certain useful identities concerning the expectation operator that I neglected to mention early on in the course. Now is as good a time as any to talk about them. Here they are:

- Suppose that  $X$  is a continuous random variable having pdf  $f(x)$ , and suppose that  $\alpha(x)$  is some function. Then,

$$\mathbb{E}(\alpha(X)) = \int_{-\infty}^{\infty} \alpha(x)f(x)dx.$$

There is also a discrete analogue of this identity.

- Suppose that  $X_1, \dots, X_k$  are random variables and that  $\lambda_1, \dots, \lambda_k$  are real numbers. Then,

$$\mathbb{E}(\lambda_1 X_1 + \dots + \lambda_k X_k) = \lambda_1 \mathbb{E}(X_1) + \dots + \lambda_k \mathbb{E}(X_k).$$

This property is called “Linearity of Expectation”.

- Suppose that  $X_1, \dots, X_k$  are independent random variables. Then, the expectation of a product is the product of expectations; that is

$$\mathbb{E}(X_1 \cdots X_k) = \mathbb{E}(X_1)\mathbb{E}(X_2) \cdots \mathbb{E}(X_k).$$

**Note:** You really do need independence here. There are examples of r.v.’s that are not independent, for which this identity fails to hold.

- Suppose that  $X$  and  $Y$  are random variables. Then,

$$\mathbb{E}_y \mathbb{E}(X|Y=y) = \mathbb{E}(X).$$

This identity is called the “Tower Property of Expectation”.

## 2 Proofs of the identities

### 2.1 Proof of the first identity

The idea of the proof is just to make a change of variable: Let  $Y = \alpha(X)$ , and let  $g(y)$  denote the pdf for  $Y$ . For the purposes of this proof (to make it simpler to describe), let us assume that  $\alpha$  is an increasing function. Then, we have that

$$\mathbb{P}(Y \leq y) = \int_{-\infty}^{\alpha^{-1}(y)} f(t)dt;$$

and so,

$$g(y) = f(\alpha^{-1}(y)) \frac{d\alpha^{-1}(y)}{dy}.$$

We then get that

$$\mathbb{E}(Y) = \int_{-\infty}^{\infty} yg(y)dy = \int_{-\infty}^{\infty} yf(\alpha^{-1}(y)) \frac{d\alpha^{-1}(y)}{dy} dy.$$

Making the substitution  $x = \alpha^{-1}(y)$  we find that

$$\mathbb{E}(Y) = \int_{-\alpha^{-1}(-\infty)}^{\alpha^{-1}(\infty)} \alpha(x)f(x)dx.$$

The proof in the discrete case is even easier. We will not bother to give it here.

### 2.2 Proof of linearity of expectation

Suppose  $f(x_1, \dots, x_k)$  is the joint pdf for  $(X_1, \dots, X_k)$ . Then, by definition,

$$\begin{aligned} \mathbb{E}(\lambda_1 X_1 + \dots + \lambda_k X_k) &= \int_{\mathbb{R}^k} (\lambda_1 x_1 + \dots + \lambda_k x_k) f(x_1, \dots, x_k) dx_1 \dots dx_k \\ &= \sum_{i=1}^k \int_{\mathbb{R}} \lambda_i x_i \left( \int_{\mathbb{R}^{k-1}} f(x_1, \dots, x_k) dx_1 \dots dx_{i-1} dx_{i+1} \dots dx_k \right) dx_i \\ &= \sum_{i=1}^k \lambda_i \int_{\mathbb{R}} x_i f_i(x_i) dx_i \\ &= \sum_{i=1}^k \lambda_i \mathbb{E}(X_i), \end{aligned}$$

where  $f_i$  denotes the marginal pdf for  $X_i$ .

The discrete case is proved similarly.

### 2.3 Expectation of a product of independent random variables

Suppose  $X_1, \dots, X_k$  are independent random variables, having joint pdf  $f(x_1, \dots, x_k)$ . Independence here implies that

$$f(x_1, \dots, x_k) = f_1(x_1) \cdots f_k(x_k),$$

where the  $f_i(x_i)$  are marginal pdfs. We have then that

$$\begin{aligned} \mathbb{E}(X_1 \cdots X_k) &= \int_{\mathbb{R}^k} x_1 \cdots x_k f(x_1, \dots, x_k) dx_1 \cdots dx_k \\ &= \int_{\mathbb{R}^k} x_1 \cdots x_k f_1(x_1) \cdots f_k(x_k) dx_1 \cdots dx_k \\ &= \left( \int_{\mathbb{R}} x_1 f_1(x_1) dx_1 \right) \cdots \left( \int_{\mathbb{R}} x_k f_k(x_k) dx_k \right) \\ &= \mathbb{E}(X_1) \cdots \mathbb{E}(X_k). \end{aligned}$$

### 2.4 Proof of the tower property of expectation

Before we prove this particular identity we need to discuss what the conditional expectation notation even means: Given a joint pdf  $f(x, y)$ , we define something called the *conditional pdf* given by

$$f(x|y) = \frac{f(x, y)}{\int_{-\infty}^{\infty} f(x, y) dx} = \frac{f(x, y)}{h(y)},$$

where  $h(y)$  is the marginal pdf for  $Y$ .

Basically, in the discrete case this would be  $\mathbb{P}(X = x | Y = y)$ ; in the continuous case we have that for a region  $A \subseteq \mathbb{R}^2$ ,

$$\mathbb{P}(X \in A | Y = y) = \int_A f(x|y) dx dy.$$

Once we have this notion, we can define conditional expectation as follows:

$$\mathbb{E}(X | Y = y) = \int_{\mathbb{R}} x f(x|y) dx.$$

And then we have that

$$\begin{aligned}\mathbb{E}_y \mathbb{E}(X|Y=y) &= \int_{-\infty}^{\infty} \mathbb{E}(X|Y=y)g(y)dy \\ &= \int_{-\infty}^{\infty} g(y) \left( \int_{\mathbb{R}} xf(x|y)dx \right) dy \\ &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} xg(y) \frac{f(x,y)}{h(y)} dx dy \\ &= \int_{-\infty}^{\infty} xg(x)dx \\ &= \mathbb{E}(X),\end{aligned}$$

where  $g(x)$  is the marginal pdf for  $X$ , given by

$$g(x) = \int_{-\infty}^{\infty} f(x,y)dy.$$